Intra- and inter-individual variation in self-reported code-switching patterns of adult multilinguals

Jean-Marc Dewaele
Birkbeck, University of London

Li Wei
Birkbeck, University of London

Corresponding author:
Dr Jean-Marc Dewaele,
Birkbeck College, University of London,
26 Russell Square,
London WC1B 5DT
UK
E-mail address: j.dewaele@bbk.ac.uk
Tel +44 207 631 6399

Abstract
The present study is a large-scale quantitative analysis of intra-individual variation (linked to type of interlocutor) and inter-individual variation (linked to multilingualism, sociobiographical variables and three personality traits) in self-reported frequency of code-switching (CS) among 2116 multilinguals. We found a significant effect of interlocutor (friends, family, colleagues and strangers) on self-reported CS. Participants who grew up and work in a multilingual environment, know many different languages, which they learnt from a young age, have advanced proficiency in various languages reported more frequent CS. Sex, Extraversion and Cognitive Empathy, but not Tolerance of Ambiguity, are linked with higher self-reported CS. We conclude that the frequency of self-reported CS depends not just on situational, complex sociobiographical and environmental factors, but it is also mediated by the personality of the multilingual.

Key words: individual differences, code-switching, multilingualism, personality

Introduction

Rodriguez-Fornells, Krämer, Lorenzo-Seva, Festman and Münte (2012) observe that: ‘Although a plethora of studies have been devoted to code-switching, an important aspect that has been neglected in psycholinguistic, linguistic, and sociolinguistic approaches to this phenomenon is the role of individual differences in language switching’ (2012, p. 1). The present study is intended to start filling this gap by investigating intra- and inter-speaker variations in code-switching (hereafter CS) patterns of adult multilinguals. We define CS as ‘changes from one language to another in the course of conversation’ (Li Wei, 2007, p. 14). There are many different structural configurations of CS (Li Wei, 2013). Whilst individuals may indeed vary the way they code-switch from one context to another, this is a topic that is beyond the scope of the present study. We focus instead on the likelihood of individual speakers to engage in CS in different settings and for different purposes.

The lack of systematic research into individual differences in CS could be linked to a number of reasons. First of all, CS can be a creative discourse strategy applied by multilingual language users in

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real-life interactions in order to achieve effective communication (Ritchie and Bhatia, 2013). This has two critical implications. One is that most studies of CS are focused on either the discourse processes whereby multilinguals combine elements from different languages to create mixed-code utterances or the constraints, linguistic and cognitive, on switching from one language to another. The study of self-reported individual differences in CS requires quantifiable and comparative data from a large number of participants. One way to collect such data is through self-reporting questionnaires, as in the case of Rodriguez-Fornells et al’s study. We will also exploit this data collection method in the present study.

The other implication is that the spontaneity of CS render the identification of CS patterns, a pre-requisite for analysing individual differences, an almost impossible task, if patterns are to be understood as regular usage shared by a group of speakers. Multilinguals are known to be very inventive, creating novel structural forms with elements of different languages spontaneously and deliberately flouting linguistic and social conventions for special effects. Rodriguez-Fornells, et al. (2012) base their study on four types of CS: switches to L1; switches to L2; contextual switches, i.e. switches that are triggered by a particular situation, topic, or environment; and unintended switches. The first two types are intended to measure switching behavior related to linguistic factors such as competence and proficiency in the target and non-target languages and semantic differences across languages. The contextual switch is designed to measure switching patterns influenced by sociolinguistic factors. And the unintentional switch is aimed to assess cognitive control and performance monitoring or problems in the control of activation of the non-target language.

In the present study, we focus exclusively on situational variation in the frequency of self-reported CS and make no assumption of either the structural configurations of the mixed-language utterances or the level of activation and cognitive control i.e. the ‘flexible and controlled manner of voluntarily guiding goal-directed behavior’, and more specifically ‘executive processes such as voluntary response inhibition and working memory that allow planned responses’ (Luna, 2009, p. 233). We believe that this is a necessary first step toward establishing individual differences in self-reported and actual CS. Details of how exactly individuals differ in combining elements of different languages to produce CS or in the level of activation and control are important topics for future investigations.

The scarcity of research on individual differences in CS could also be due to the confusion over what the very notion of individual differences means to different researchers. For instance, does it refer to the same individual’s different usage in different contexts, i.e. intra-speaker variation, or is it differences between speakers in producing the same code-switching pattern, i.e. inter-speaker variation, or both? Moreover, there is huge diversity in the factors researchers choose to use in studying individual differences. Language proficiency and history of language acquisition and learning are probably the most frequently used factors. But other factors, such as personality, sex, and interpersonal relationship, which are the core factors in psychological research on individual differences, remain under-explored in CS research. The present study will examine both intra- and inter-speaker variation with a particular focus on type of interlocutor, prior and current linguistic practices, age, sex, education and personality traits such as extraversion, tolerance of ambiguity, and cognitive empathy.

The paper is structured as follows. We first review the existing research on individual variation in CS, focusing on sociolinguistic studies. We then introduce the personality traits which we hypothesise to have a prima facie link with self-reported CS. Our research instruments include an online questionnaire specially designed for the present study, the short version of the Eysenck Personality Questionnaire (EPQr) (Eysenck, Eysenck & Barrett, 1985); the adapted measure of Cognitive Empathy (Baron-Cohen and Wheelwright, 2004); and the adapted Tolerance for Ambiguity Scale (TAS) (Herman, Stevens, Bird, Mendenhall & Oddou, 2010). We will describe the design of our empirical study, and our research questions and hypotheses. Subsequently, we will test four specific research hypotheses with our empirical data. The implications of our findings are discussed in the concluding section.
Existing sociolinguistic studies of individual variations in CS

Multilingual language users code-switch for a variety of reasons that sociolinguists have termed as ‘motivations’. Gumperz (1982) and Blom and Gumperz (1972) introduced the notions of situational versus metaphorical CS. Situational CS is triggered by a change of address, topic or setting, whereas metaphorical CS does not involve any change of these variables but are motivated by communicative functions such as quotation, addressee specification, interjection, reiteration, message qualification, or personalization versus objectivation; in other words, special communicative effects. To these we can add i) cultural specificity - certain notions or concepts may only exist in one language or are simply better expressed in a specific language; ii) experiential association – as Grosjean says: ‘Bilinguals usually acquire and use their languages for different purposes, in different domains of life, with different people. Different aspects of life often require different languages’ (2010, p. 29). And iii) identity and identification - CS can be strategically employed when the speaker wishes to show involvement, group membership, expertise, power, and also excluding someone who does not know the language. Myers-Scotton’s (1993) investigation of CS in conversations of multilingual Africans (mostly Kenyans living in Nairobi) considered different types of CS, its structure, and some social motivations. She concluded that CS is a type of skilled performance, rather than a means of overcoming linguistic difficulties. CS expressed social import and could be interpreted as identity negotiations (p. 151). The work on social motivations for CS, however, does not generally discuss individual variations.

Since the 1980s, there has been a group of sociolinguists who have attempted to explore the question whether different groups of bilingual and multilingual speakers would use CS in different ways in different contexts. Working in the Labovian variationist sociolinguistics paradigm, Poplack (1980), found that speakers in the East Harlem Spanish-English bilingual community in New York City adapted their language choice between Spanish and English and their CS patterns according to the situation they found themselves to be in, e.g. home, school and neighbourhood. On the whole, the most significant variation existed between individual speakers, especially between children who were receiving bilingual schooling or monolingual schooling, with the former switching much more. Gardner-Chloros (1991), using a similar approach, surveyed the use of French and Alsatian, and also switching between the two, in Strasbourg’s department stores. Complex language choice and CS patterns were revealed. For example, the highest rate of CS was found among the youngest group of shoppers, who were also the most French-speaking, when they were in the most Alsatian-speaking store, whereas the same group of speakers used little Alsatian, and little CS also, in another, more linguistically neutral store. Gardner-Chloros explained the variation in terms of speech accommodation and language prestige. Li Wei (1995) examined a number of CS patterns - inter-speaker, inter-sentential inter-clausal, and content words – by three groups of Chinese-English bilinguals: students, British-born Chinese youth and long-term immigrants from Hong Kong, all living in the North East of England. He found that variations existed across these groups according to the bilingual experience and social network contacts the speakers had, as well as the topic, setting and interlocutor of the interaction.

Ritchie and Bhatia (2013) noted that CS is linked to social roles and relationships between participants. Other factors include situational factors, message-intrinsic considerations and finally language attitudes including social dominance and security (p. 378).

Dewaele (2010) investigated the effect of type of interlocutor and conversation topic on self-reported frequency of CS among 1453 multilinguals who filled out the BEQ (Bilingualism and Emotions Questionnaire - Dewaele & Pavlenko, 2001-2003) and among 20 multilinguals who were interviewed about their CS practices. Self-reported CS was found to be significantly lower when speaking about neutral topics compared to personal or emotional topics (Dewaele, 2010, pp. 196-197). When strong emotions need to be verbalized quickly, and the speaker realizes that it would take too much time to express them in the weaker output language, possibly with unwanted effects, CS might seem like an acceptable option (p. 213). The analysis of the narratives showed that CS is strongly linked to the perceived emotionality of the languages and that CS is deployed strategically. The typical direction of
the CS in situations where strong emotion had to be expressed was from the foreign language to the L1. Some participants, often of Arab or Asian origin, reported CS in the opposite direction, especially for the expression of anger and swearing. They said that for them, CS to English (a foreign language) allowed them to escape the social constraint that weighs on them in their home environments, where anger cannot be displayed as openly as in English and where swearing carries strong social stigma (p. 212).

Dewaele also found that the higher frequency of CS with known rather than less known, or unknown, interlocutors was linked to a conscious, strategic choice of the speaker. In dealing with an unknown interlocutor the most logical option is to stick to the language in which the interaction was started. Conversations with known interlocutors allow CS to the shared languages (p. 219). Resnik (2012) reported similar effects of the type of interlocutor on self-reported CS in emotional interactions with German, Chinese and Japanese foreign language users of English.

Some significant sex effects emerged in Dewaele (2010), with female participants reporting more frequent CS with friends and on more emotional topics, while male participants reported more CS with strangers. These sex effects seem rather unusual considering Gardner-Chloros (2009) who states that CS is unconnected with sex, although ‘it intersects with a large number of intervening variables which are themselves connected with gender issues’ (2009, p. 82).

Another extralinguistic variable found to be associated with the frequency of CS is the linguistic environment in which an individual grows up or is currently living. Family and community norms may affect the extent of CS of the individual (Lanza, 2008).

One factor that has hitherto not been examined systematically is personality. Gardner-Chloros (2008) points out that the way in which multilinguals use CS is part of the way in which they present themselves as speakers, i.e. ‘styling the self’. She further observes that among the many independent variables which have been studied in relation to CS, personality variables affecting the amount and type of CS have on the whole been omitted.

In the present study, we hypothesise that certain personality traits contribute to a higher self-reported frequency of CS, possibly in interaction with particular social variables such as context and language background. In the following section, we discuss the key personality traits that we intend to investigate.

**Personality traits**

Psychologists generally agree that personality traits are hierarchically organized with three or five broad, independent dimensions at the top - and a larger number of more specific traits lower down (Cervone & Pervin, 2013). Personality questionnaires allow researchers to establish profiles of participants. In the present study, we focus on one dimension out of the so-called ‘Giant three’ Extraversion versus Introversion, and two ‘lower-order’ traits, Cognitive Empathy and Tolerance of Ambiguity. All these three traits have been found to be associated with second language learning and multilingualism (Dewaele, 2012; Dewaele & Li Wei, 2012, 2013). We hypothesize that they are also linked to CS.

**Extraversion - Introversion**

Eysenck and Eysenck (1985) argue that variation on the Extraversion – Introversion dimension is linked to the amount of cortical arousal, which leads to different behavior, also different communicative behaviour in a foreign language (Dewaele, 2012). While extraverts are under-aroused, introverts are over-aroused. This has behavioral consequences: extraverts compensate for their suboptimal arousal levels by tending towards activities that involve greater sensory stimulation while introverts will tend to avoid over-arousing situations. While the typical extravert is gregarious, ‘craves excitement, takes chances, often sticks his neck out, acts on the spur of the moment’ (Eysenck & Eysenck 1964, p. 8), the typical introvert ‘is a quiet, retiring sort of person, introspective, (...) tends to plan ahead, ‘looks before he leaps,’ and (...) does not like excitement’ (p. 8). The extraverts’ inclination to take risks seems to extend to their linguistic behavior. Extravert L2 users use more stigmatised
variants than their more introverted peers and more likely to talk about emotion, a particularly
challenging topic in a foreign language (Dewaele 2004; Dewaele & Pavlenko, 2002). Li, Chen and
Xiao (2009) have argued that extraverts possess superior pragmatic competence and awareness, which
allow them to engage in more risky linguistic behaviour. In his overview of individual differences and
creativity, Kaufman (2011, p. 681) reports that extraversion has been linked to creativity in some
domains but not in others. The study of Peterson, Smith and Carson (2002) found a significant positive
correlation between extraversion and creativity (p. 1143). King, Walker and Broyles (1996) also found
a significant positive correlation between extraversion and verbal creativity. It is thus plausible that
extraverts are more likely to engage in creative CS.

**Cognitive empathy**

Empathy has been defined as ‘the ‘glue’ of the social world’ (Baron-Cohen & Wheelright, 2004, p.
193). It refers to the ability to empathise, to understanding what other people might be thinking or how
they might be feeling (p. 193). People with high levels of empathy are better at understanding the
intentions of others, are more accurate in predicting their behaviour, and are better able to recognise the
emotion of their interlocutor. Empathy is thus crucial in social interactions, ‘drawing us to help others
and stopping us from hurting others’ (p. 193).

Dewaele and Li Wei (2012) investigated the relationship between multilingualism and cognitive
empathy among 2158 mono- and multilinguals (who also contributed the data for the present study). A
significant positive correlation emerged between multilingualism (operationalised as advanced levels
of proficiency in several foreign languages and frequent use of these languages) and cognitive empathy.
They concluded that intense multilingual practice makes multilinguals more skilful in conversations as
they learn to see the world from their interlocutor’s point of view. It is possible that the ability to
empathise with a bilingual or multilingual interlocutor might be linked to increased CS, as this is a way
of highlighting the specific links between the speaker and the interlocutor. A high level of empathy
might nudge a bilingual or multilingual to converge to a higher frequency of CS to reflect the pattern of
the interlocutor.

**Tolerance of Ambiguity**

According to Furnham and Ribchester (1995, p. 179), ‘The person with low tolerance of ambiguity
experiences stress, reacts prematurely, and avoids ambiguous stimuli. At the other extreme of the scale,
however, a person with high tolerance of ambiguity perceived ambiguous situations/stimuli as
desirable, challenging, and interesting and neither denies nor distorts their complexity of incongruity’.

Dewaele and Li Wei (2013a), using the database on which the present study is based, found that
monolinguals and bilinguals scored significantly lower on Tolerance of Ambiguity (TA) compared to
multilinguals. Moreover, participants with higher levels of multilingualism, as measured through total
proficiency in multiple languages, also scored significantly higher on TA. A multilingual upbringing
had no effect on TA but those who have lived abroad scored significantly higher on TA. The authors
concluded that TA is determined by individuals’ social-linguistic-cultural environment and especially
by the experience of having to survive in a foreign cultural and linguistic environment. A prolonged
stay in a new environment requires a sustained and conscious effort to acquire the new sociopragmatic
rules and cultural values. CS could be considered as a way to make the interaction more interesting. We
will consider the potential link between TA and CS.

**Research questions and hypotheses**

The present study aims to address the following research questions:

1) Does the amount of self-reported CS vary according to the interlocutor type? We
   expect most frequent CS with friends, followed by colleagues, family, and least of all with
   strangers.
2) Are prior and current linguistic practices linked to the amount of self-reported CS? We expect participants who grew up in a linguistic and ethnic diverse environment to code-switch more. We also expect participants who know more languages, who have advanced knowledge of several languages, who grew up with two or more languages before age 3, who learned languages early in life to code-switch more. Finally, we expect participants who work in multilingual environments to code-switch more.

3) Are sociobiographical variables linked to the amount of self-reported CS? We expect female participants to code-switch more with known interlocutors. Age and education level could also be linked to self-reported CS.

4) Are personality traits linked to the amount of self-reported CS? We expect participants who score high on Extraversion, Tolerance of Ambiguity, Cognitive Empathy to code-switch more.

Method

Participants

A total of 2116 multilinguals (1564 females, 443 males) filled out the online questionnaire on the use of CS that has been designed specially for the present study. The mean age was 34.5 yrs (SD = 12.1). A closer look showed that 90 participants were in their teens, 799 were in their twenties, 583 in their thirties, 347 in the forties, 190 in their fifties, 87 in their sixties, 8 in their seventies and one participant aged 84. Participants were generally highly educated with only 31 having a high school diploma, 650 a Bachelor’s degree, 772 a Master’s degree, and 657 a PhD. This majority of highly educated, female participants is quite typical in web-based language questionnaires (Wilson & Dewaele, 2010).

The participants reported 204 different nationalities, including many participants with double nationalities. The largest group came from the USA (n = 478), followed by British (n = 299); Dutch (n = 145); Belgians (n = 81), Germans (n = 81), Canadians (n = 76), Polish (n = 65), French (n = 58), Spaniards (n = 42), Chinese (n = 41). English was the most frequent L1 (n = 843), followed by Dutch (n = 195), French (n = 155), Spanish (n = 138), German (n = 124). The most frequent L2 was English (n = 924) followed by French (n = 455), and Spanish (n = 248). The pattern was different for the L3 with French coming first (n = 422), followed by German (n = 330) and English (n = 248). The most frequent L4s were German (n = 205), Spanish (n = 196) and French (n = 174). The most frequent L5 was Spanish (n = 101), followed by Italian (n = 69) and French (n = 50). Mean age of acquisition of the L2 was 10.1 years (SD = 5.4), this increased to 15 years for the L3 (SD = 6.4), 18.3 years for the L4 (SD = 7.8) and 21.7 for the L5 (SD = 8.6).

We used participants’ information on self-perceived proficiency in their various languages to develop a global measure of multilingualism, first presented in Dewaele and Stavans (2012). The ‘total proficiency score’ is the sum of self-perceived proficiency scores collected on 5-point Likert scales for oral proficiency and written proficiency in up to 6 languages (including 2 L1s) (maximal possible score 10 X 6 = 60). Dewaele and Stavans (2012) argued that such a measure is potentially useful to distinguish pentalinguals with limited knowledge of three languages from trilinguals with advanced knowledge of 3 languages. The trilingual might know fewer languages, but knowing them better makes that individual more strongly multilingual. In other words, we avoid the lack of clarity inherent to labels such as ‘bilingual, trilingual, quadrilingual etc’, where every language is included, despite the fact that knowledge in some can be very limited. In the present sample total proficiency scores vary between 5 and 55 with a mean of 25.5 (SD = 8.0). Participants with scores that were more than 1 standard deviation below the mean were categorised as ‘Low Proficiency’ (n = 244), those with scores that were more than 1 standard deviation above the mean were categorised as ‘High Proficiency’ (n = 278), while the remaining participants were categorised as ‘Medium Proficiency’ (n = 1510). We feel also that it makes sense to look for a link between this global language measure and CS, rather than proficiency measures for particular languages (L1, L2, L3...).
The sample consists of 399 bilinguals, 566 trilinguals, 557 quadrilinguals, 359 pentalinguals, 143 sextalinguals, 54 septalinguals, 21 octalinguals, 9 nonalinguals, 5 participants knew 10 languages, and 1 participant reported 12 languages. A single category was created including all participants with 6 or more languages. A majority of participants (n = 1825) reported having been exposed to a single L1 before age 3, a minority reported growing up with two L1s (n = 274) and 17 participants grew up with three languages (the latter two groups were aggregated).

Mean score for ethnic and linguistic diversity during the participants’ childhood was rather low (M = 1.26, SD = 1.23, on a 4-point Likert scale). However, the mean score for linguistic diversity in the participants’ workplace was higher (M = 2.59, SD = 1.23, on a 4-point Likert scale).

**Instruments**

The data was collected through the snowballing technique. The anonymous online CS questionnaire was an open-access survey, advertised through several listservs, targeted emails to teachers and students, and informal contacts asking them to forward the link to friends. The introductory paragraph stated: ‘The aim of this research is to help us better understand the link between personality and linguistic behavior of people knowing one or several languages’. It remained online over a four month period and attracted responses from mono- and multilinguals across the world. We removed the data provided by the 41 monolinguals for the present study, as they did not code-switch between languages. Because participants left occasional questions blank, totals for specific variables can vary. Participants also filled out a short sociobiographical questionnaire with questions about sex, age, education, language history and present language use. The research design and questionnaires received ethical clearance from the research institution.

Buchanan (2007) observes that the use of online questionnaires in psychological research is increasing exponentially. The main advantage is of an ‘economic’ nature: large amounts of data can be collected automatically at a fraction of the cost and time of ‘pen and paper’ equivalents (Buchanan, 2007). This allows researchers to reach larger and more diverse samples from all over the world. Internet samples have been found to be more diverse in terms of sex, age, race, socio-economic status and geographical location than the pen-and-paper samples (Gosling et al., 2004). Scores of personality dimensions and patterns of socializing are comparable in internet samples and pen-and-paper samples (Gosling et al., 2004). Moreover, online versions of traditional questionnaires do not compromise the psychometric properties of such measures (Denissen, Neumann & van Zalk, 2010).

Wilson and Dewaele (2010) have pointed out that in multilingualism research participants must meet specific linguistic criteria, have high levels of metalinguistic and metacognitive awareness, and must be able and willing to engage with the questions on language preferences and use (2010, p. 108). We feel that multilinguals are perfectly aware of the amount of CS they use in specific interactions, and they have no reason to lie about the frequency with which they use CS, as they perceive the phenomenon as something positive (Dewaele & Li Wei, 2013b). The reliability of the data might in fact be stronger in linguistic internet-based research, as anonymous volunteers would not benefit in any way from falsifying answers (p. 108). Another advantage of internet-based questionnaires is that they reduce social desirability (the tendency of participants to answer questions in a manner that they imagine will be viewed favourably by the researcher), which leads to increased levels of honesty (and therefore higher validity in the case of self-report) (Joinson, Paine, Buchanan & Reips, 2008). Finally, a sample of more than 2000 multilinguals from all over the world ensures strong ecological validity, as the effects of local social, political and historical factors linked to particular languages or linguistic practices are evened out (Wilson & Dewaele, 2010).

In addition to the background questionnaire, participants filled out the items for the personality dimension Extraversion extracted from the EPQr (Eysenck, Eysenck & Barrett, 1985), also based on self-reported behaviour. Participants filling out the EPQr are invited to tick either ‘yes’ or ‘no’ for 12 items for each dimension. One item for Extraversion is for example: ‘Are you a talkative person?’.

Barrett, Petrides, Eysenck and Eysenck (1998) demonstrated the factorial similarity of the dimensions in data collected from 34 countries, which suggests that the Eysenck factors are strongly
replicable across the world. The EPQr is considered robust (Barrett, 1999). Mean scores for extraversion was as follows: \( N = 1888, M = 7.6, SD = 3.4 \). Internal consistency, as measured by Cronbach alpha coefficient, was high (0.84). Participants with scores that were more than 1 standard deviation below the mean were categorised as ‘Introverts’. Those with scores ranging from 1 standard deviation below the mean to 1 standard deviation above the mean were categorised as ‘Ambiverts’. Finally, those with scores that were more than 1 standard deviation above the mean were categorised as ‘Extraverts’. As a result we have 399 Introverts, 1008 Ambiverts and 481 Extraverts.

Participants also filled out the Tolerance for Ambiguity Scale (TAS), a 12–item questionnaire with 5-point Likert scales, contextualized to cross-cultural contexts (Herman et al., 2010, p. 60). According to the authors, it ‘can be used in cross-cultural research and practice to assess individual TA’ (p. 62). It consists of four distinct dimensions: (1) valuing diverse others; (2) change; (3) challenging perspectives; (4) unfamiliarity (p. 62). We used 11 items of out the original 12-item TAS scale and made some minor stylistic adaptations. The items had to be rated on a 5-point Likert scale. A Cronbach alpha analysis revealed modest but sufficient internal consistency reliability: 0.64. Mean score for TA is 28.9, \( SD = 5.6 \), with a range between 4 and 44. Participants were also categorised in three groups (low TA, medium TA, high TA), following the procedure set out for extraversion. As a result 277 participants were labelled ‘low TA’, 1337 ‘medium TA’ and 281 ‘high TA’.

Finally, participants filled out an adapted version of Baron-Cohen and Wheelwright’s (2004) questionnaire to measure Cognitive Empathy (CE). We selected five items with the highest factor loadings (Lawrence et al., 2004, p. 915) on the dimension of CE, which ‘measures the appreciation of affective states’ (p. 918). One such item was ‘I am good at predicting how someone will feel’. The items had to be rated on a 5-point Likert scale. This subscale was internally consistent (Cronbach's Alpha = 0.84)\textsuperscript{vii}. Mean score of CE is 18.5, \( SD = 3.5 \), with a range between 5 and 25. Three groups were created (low CE, medium CE, high CE), following the procedure set out for Extraversion and TA. As a result 230 participants were labelled ‘low CE’, 1397 were in the middle group (‘medium CE’) and 267 were labelled ‘high CE’.

Data about CS practices of the participants were elicited through the following general closed question: ‘Do you switch between languages within a conversation with certain people?’ It then specified four situations: ‘when speaking with some friends’, ‘with some family members’, ‘with strangers’ and ‘with some colleagues or clients at work/school’\textsuperscript{viii}. Participants were asked to choose a response on a 5-point Likert scale, i.e. never, rarely, sometimes, frequently, always. We are therefore more concerned with the frequency of CS with different interlocutors in different contexts rather than with the structural patterns of switching.

A series of one-sample Kolmogorov-Smirnov tests showed that the values for self-reported frequency of CS with the four types of interlocutors are not normally distributed (Kolmogorov-Smirnov Z-values vary between 8.2 and 10.4, all significant at \( p < .0001 \)). More than a third of participants reported never using CS with family members nor strangers. The proportion of non-CS users dropped to 22% in interactions with colleagues and 12% in interactions with friends where the frequency category ‘sometimes’ is the largest (a third of participants). The complete distribution is presented in Figure 1.

Figure 1. Distribution of frequency of self-reported CS with four types of interlocutors
As a consequence, we have used non-parametric statistical techniques\textsuperscript{viii}.

**Results**

A Friedman’s ANOVA test for related samples revealed that the effect of type of interlocutor is highly significant on the amount of self-reported CS ($N = 1997$, $Chi^2 = 1047.7$, $df = 3$, $p = .0001$) (Table 1). This result confirms that multilinguals report CS most with their friends and least with strangers. The standard deviation is also the smallest for these two situations, especially for strangers. There is much more dispersion around the mean for self-reported CS with family members. This suggests that there is more variation in CS practices within families, compared to the groups of friends, colleagues or strangers. This result is somewhat surprising as family members are likely to share similar language backgrounds.

**Table 1. Mean scores on self-reported frequency of CS**

<table>
<thead>
<tr>
<th>Code-switching with...</th>
<th>M</th>
<th>SD</th>
</tr>
</thead>
<tbody>
<tr>
<td>Friends</td>
<td>2.16</td>
<td>1.15</td>
</tr>
<tr>
<td>Colleagues</td>
<td>1.70</td>
<td>1.19</td>
</tr>
<tr>
<td>Family</td>
<td>1.46</td>
<td>1.41</td>
</tr>
<tr>
<td>Strangers</td>
<td>1.09</td>
<td>1.00</td>
</tr>
</tbody>
</table>

The next cluster of variables linked to participants’ linguistic history and current practice show the extent to which it is linked to self-reported CS. The Kruskal-Wallis tests showed that participants who grew up in an ethnically and linguistically diverse environment report more frequent CS in all situations except with strangers (Table 2 and Figure 2).

**Table 2. The effect of linguistic history and current practice on self-reported frequency of CS (Kruskal-Wallis Chi$^2$ values)**

<table>
<thead>
<tr>
<th></th>
<th>Friends</th>
<th>Family</th>
<th>Strangers</th>
<th>Colleagues</th>
</tr>
</thead>
<tbody>
<tr>
<td>Linguistic diversity</td>
<td>10.7*</td>
<td>29.5***</td>
<td>4.3</td>
<td>8.3*</td>
</tr>
<tr>
<td>during childhood</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Number of languages known</td>
<td>33.9***</td>
<td>51.6***</td>
<td>38.8***</td>
<td>35.6***</td>
</tr>
</tbody>
</table>
Total Proficiency was found to have a significant effect on self-reported frequency of CS in all situations (Table 2). The High Total Proficiency group reported more frequent CS with all interlocutors (Figure 4).
A Mann-Whitney test showed significant differences between early and later bi- and multilinguals in interactions with friends and family, with a marginal difference in interactions with strangers (Table 3). As expected early bi- and multilinguals reported using more CS than later bi- and multilinguals (Figure 5).

Table 3. *The effect of early bi- or multilingualism on self-reported frequency of CS (Mann-Whitney)*

<table>
<thead>
<tr>
<th>Interlocutor</th>
<th>Z</th>
</tr>
</thead>
<tbody>
<tr>
<td>Friends</td>
<td>-4.3***</td>
</tr>
<tr>
<td>Family</td>
<td>-11.9***</td>
</tr>
<tr>
<td>Strangers</td>
<td>-1.8</td>
</tr>
<tr>
<td>Colleagues</td>
<td>-0.7</td>
</tr>
</tbody>
</table>

*** p < .0001
A Spearman Rank correlation analysis revealed significant negative values between Age of Onset of Acquisition (AoA) in the L2 and L3 and self-reported CS with family, friends and strangers (but not with colleagues). This suggests that multilinguals who started learning their L2 and L3 earlier in life reported using more CS in specific situations. The AoA of the L4 and L5 (both over the age of 18) was unrelated to self-reported CS frequency (see Table 4).

Table 4. The relationship between Age of onset of Acquisition and CS (Spearman Rho)

<table>
<thead>
<tr>
<th>Language</th>
<th>Measure</th>
<th>Friends</th>
<th>Family</th>
<th>Strangers</th>
<th>Colleagues</th>
</tr>
</thead>
<tbody>
<tr>
<td>L2</td>
<td>Rho</td>
<td>-.124***</td>
<td>-.211***</td>
<td>-.052*</td>
<td>-.033</td>
</tr>
<tr>
<td></td>
<td>N</td>
<td>2028</td>
<td>2012</td>
<td>2011</td>
<td>2009</td>
</tr>
<tr>
<td>L3</td>
<td>Rho</td>
<td>-.060*</td>
<td>-.165***</td>
<td>-.044</td>
<td>-.022</td>
</tr>
<tr>
<td></td>
<td>N</td>
<td>1633</td>
<td>1622</td>
<td>1623</td>
<td>1618</td>
</tr>
<tr>
<td>L4</td>
<td>Rho</td>
<td>-.057</td>
<td>-.098**</td>
<td>-.044</td>
<td>-.047</td>
</tr>
<tr>
<td></td>
<td>N</td>
<td>1062</td>
<td>1061</td>
<td>1058</td>
<td>1059</td>
</tr>
<tr>
<td>L5</td>
<td>Rho</td>
<td>-.039</td>
<td>-.012</td>
<td>-.007</td>
<td>-.020</td>
</tr>
<tr>
<td></td>
<td>N</td>
<td>514</td>
<td>512</td>
<td>511</td>
<td>511</td>
</tr>
</tbody>
</table>

• p < .05, ** p < .001, *** p < .0001

The degree of multilingualism in the work environment of participants also has a significant effect on self-reported frequency of CS in all situations except in the family sphere (Table 2). Participants in more multilingual workplaces report more frequent CS with colleagues and clients (Figure 6).
The next research question deals with the effect of sociobiographical variables such as age, sex and education level. A Spearman correlation analysis showed that age is positively linked to self-reported frequency of CS in interactions with family and strangers, and negatively linked to frequency of CS in interactions with friends and colleagues. In other words, older participants report using CS more with family members and strangers and less with friends and colleagues (Table 5).

Table 5. The relationship between age and self-reported frequency of CS (Spearman Rho)

<table>
<thead>
<tr>
<th>Interlocutor</th>
<th>N</th>
<th>Spearman Rho</th>
</tr>
</thead>
<tbody>
<tr>
<td>Friends</td>
<td>2022</td>
<td>-.071**</td>
</tr>
<tr>
<td>Family</td>
<td>2006</td>
<td>.133***</td>
</tr>
<tr>
<td>Strangers</td>
<td>2006</td>
<td>.081***</td>
</tr>
<tr>
<td>Colleagues</td>
<td>2003</td>
<td>-.043</td>
</tr>
</tbody>
</table>

* p < .05, ** p < .001, *** p < .0001

The Kruskal-Wallis analyses reveal significant differences in self-reported CS between the various education levels for two types of interlocutors (Table 6). Participants with higher levels of education reported more CS with colleagues and family members.

A Mann-Whitney test shows significant differences between male and female participants in interactions with friends, family and colleagues (Table 6). Female participants report more CS in these interactions than the males.

Table 6. The effect of education level and sex on self-reported frequency of CS (Kruskal-Wallis Chi² and Mann-Whitney Z)

<table>
<thead>
<tr>
<th>Interlocutor</th>
<th>A-level</th>
<th>BA</th>
<th>MA</th>
<th>PhD</th>
<th>Chi²</th>
<th>Female</th>
<th>Male</th>
<th>Z</th>
</tr>
</thead>
<tbody>
<tr>
<td>Friends</td>
<td>2.0</td>
<td>2.2</td>
<td>2.2</td>
<td>2.2</td>
<td>1.1</td>
<td>2.2</td>
<td>1.9</td>
<td>-4.4**</td>
</tr>
<tr>
<td>Family</td>
<td>1.5</td>
<td>1.4</td>
<td>1.4</td>
<td>1.6</td>
<td>8.6*</td>
<td>1.5</td>
<td>1.2</td>
<td>-4.6**</td>
</tr>
<tr>
<td>Strangers</td>
<td>1.1</td>
<td>1.1</td>
<td>1.1</td>
<td>0.1</td>
<td>1.1</td>
<td>1.1</td>
<td>1.1</td>
<td>-0.3</td>
</tr>
<tr>
<td>Colleagues</td>
<td>1.2</td>
<td>1.7</td>
<td>1.6</td>
<td>1.8</td>
<td>10.2*</td>
<td>1.7</td>
<td>1.6</td>
<td>-2.5*</td>
</tr>
</tbody>
</table>

* p < .05, ** p < .001

The final research question focuses on the effect of some personality traits on self-reported CS.
The Kruskal-Wallis tests did reveal a highly significant effect of Extraversion on self-reported CS in four situations (Table 7).

Table 7 The effect of personality variables on self-reported frequency of CS (Kruskal-Wallis \( \chi^2 \) values)

<table>
<thead>
<tr>
<th>Trait</th>
<th>Friends</th>
<th>Family</th>
<th>Strangers</th>
<th>Colleagues</th>
</tr>
</thead>
<tbody>
<tr>
<td>Extraversion</td>
<td>32.2***</td>
<td>13.7**</td>
<td>20.0**</td>
<td>21.6**</td>
</tr>
<tr>
<td>Tolerance of Ambiguity</td>
<td>2.53</td>
<td>2.0</td>
<td>2.7</td>
<td>0.8</td>
</tr>
<tr>
<td>Cognitive Empathy</td>
<td>14.0**</td>
<td>5.5</td>
<td>0.3</td>
<td>4.9</td>
</tr>
</tbody>
</table>

* p < .05, ** p < .001, *** p < .0001

A look at the means shows that more extraverted participants report more CS with the various interlocutors (Figure 7).

Figure 7. Extraversion and self-reported frequency of CS

Tolerance of Ambiguity has no effect on self-reported frequency of CS with the various interlocutors (Table 7). Cognitive Empathy has a significant effect on self-reported frequency of CS with friends and a marginal effect on CS with family members and colleagues or clients (Table 7). A look at the means shows that the High CE group reports more frequent CS with friends (Figure 8). The same (but only marginally significant) pattern exists for self-reported frequency of CS with family members and colleagues or clients (Figure 8).
Figure 8. Cognitive Empathy and self-reported frequency of CS

Discussion

The analyses of the data show that our first hypothesis was confirmed: self-reported CS was most frequent in interactions with friends, followed by interactions with colleagues and family members, and it was least frequent in interactions with strangers.

Our second hypothesis was also confirmed: participants’ linguistic history - including the environment in which they grew up, current linguistic practices and linguistic diversity in their current workplace all have a significant effect on self-reported CS.

Our third hypothesis was partly confirmed as participants’ age was correlated with frequency of self-reported CS (though in different directions). Sex was significantly linked to self-reported CS: female participants reported more CS in interactions with friends, family and colleagues. The effect of education was significant in interactions with family members and colleagues where highly educated participants reported more CS.

Our final hypothesis was partly confirmed: the degree of Extraversion and Cognitive Empathy was significantly linked to higher levels of self-reported CS, though only with friends for CE. However, Tolerance of Ambiguity had no effect on self-reported CS.

These findings show that multilinguals adapt the frequency of CS to the type of interlocutor they are interacting with. Just as in Dewaele (2010) we found that self-reported CS is less frequent with unknown interlocutors, probably because speakers need to know which languages they have in common with their interlocutors before starting CS. Unsurprisingly, when the interlocutor is known to the speaker, there is a higher likelihood of CS. The highest amount of CS was reported with friends, followed by colleagues. The amount of CS in interactions with family members is slightly lower. One possible reason for this is that the groups of friends and colleagues are probably more heterogeneous in linguistic terms, whereas as family members may form a more homogeneous group, with clear language preferences. Rodriguez-Fornells (personal communication) suggested that the fact that more CS is reported with friends than with family members might be a simple effect of sample bias, where this population comes from, and how their familial environment was. This result might simply reflect the fact that the participants are now not too much exposed to their families anymore. It is equally
possible that CS was banned within the family home of some participants in order to maintain the minority language. However, a closer investigation confirmed that the early bi- and multilinguals reported significantly more CS with their family members and friends compared to participants who became bi- and multilingual after the age of 3.

The fact that participants who had grown up in highly multilingual and ethnically diverse environments reported using more CS with known interlocutors shows that linguistic practices in childhood continue to resonate through a multilingual’s life. Less surprising is the finding of a positive link between multilingualism in the work environment and self-reported CS (except with family). Indeed, one can assume that our participants conformed to the unwritten rules of their community of practices, which involve CS. However, they did not follow these rules when interacting with family members – which probably constitutes a different community of practice.

The relationship between age and self-reported CS was unexpected because it was positive with some interlocutors, and negative with others. There is no obvious reason why older participants would use more CS more with family members and strangers but on the other hand, use significantly less CS with friends and colleagues. These results differ from those reported by Dewaele (2010, p. 199), namely an absence of correlation between age and self-reported CS in interactions with friends and strangers, and a significant positive correlation between self-reported CS when talking to colleagues or speaking in public. The difference cannot be linked to a different age profile of the sample, as the mean age in Dewaele (2010) was 34 years, identical to the one in the present study. Had the link between age and CS with different interlocutors been consistent, it could have been associated with changes in cognitive control, which reaches a peak in adulthood and remains stable before declining in old age (Luna, 2009).

The significant sex effect in self-reported use of CS in interactions with friends, family and colleagues partly confirms previous research. Dewaele (2010, p. 198) found that the female participants reported using CS significantly more with friends, but not with colleagues. Male participants were also found to report more CS in interactions with strangers, a pattern that is absent in the present study. It is difficult to speculate on the reason why our female multilinguals report more CS with friends, family and colleagues. Are they conforming to rules of their speech communities? Does it reflect a greater involvement of our female participants in conversations (cf. Tannen, 1993)?

Education level was found to have significant effect on self-reported CS with two types of interlocutor: colleagues and family members. It is again hard to speculate on a possible cause. That highly educated participants report using more CS with colleagues could be linked to the type of work environment they function in. Indeed, a Pearson Chi² analysis showed that highly educated participants were significantly more likely to work in highly multilingual workplaces (Pearson Chi² = 44.7, df = 12, p < 0.0001). It is equally possible that the families of the highly educated participants were more multilingual, but we have no information on the degree of multilingualism of the participants’ families.

One of the most striking findings of the present study is the link between two personality traits and self-reported CS. The positive link between Extraversion and self-reported CS can be understood if we see CS as a form of impulsive linguistic risk-taking, or verbal acrobatics, giving the speaker a certain thrill. More introverted speakers report engaging less frequently in this verbal behaviour. The more gregarious multilingual extraverts might be more inclined to use CS once they discover they share another language with an interlocutor. By doing so, they know that they underlie the common ground with the interlocutor and converge towards them. The relationship between extraversion and CS could also be linked to the higher levels of (verbal) creativity of extraverts (King et al., 1996; Peterson et al., 2002). Metaphorically one could compare a monolingual interaction to a monochrome exchange and a multilingual interaction to something more colourful. Creative multilinguals can insert threads in different colours in the exchange and hence insert an element of novelty, uniqueness and surprise.

Tolerance of Ambiguity was unrelated to frequency of CS. This is surprising because we had assumed that CS is a way of introducing and sharing ambiguity in the interaction. Indeed, the pupils described in Li Wei and Chao-Jung Wu (2009) deliberately used CS to create ambiguity and undermine the monolingual teacher’s authority. While CS can serve this function, the opposite might be true in
general, namely that CS allows multilinguals to avoid ambiguity. By switching to another shared language, communication is more efficient and to the point, it allows the multilingual to find ‘le mot juste’ (Grosjean, 2010, p. 53). One could even argue that multilinguals that share a language and can fluently CS may find this speech style less ambiguous than monolingual-mode. The possibility of substituting a word in the other language if necessary, or use a colloquial saying with the interlocutor may introduce greater security into the interaction. Whilst our study does not aim to investigate the cognitive control in CS, one may speculate, based on existing evidence that links language switching with executive control (e.g. Rodriguez-Fornells, De Diego Balaguér & Münte, 2006; Soveri et al., 2011), that CS requires more selective attention; therefore, it would be associated with precision rather than ambiguity.

The positive link between Cognitive Empathy and self-reported CS in interactions with friends suggests that multilinguals might resort to CS if they feel it could help their friend. In other words, their ability to see the friend’s perspective may push them to CS if they feel it could help the flow of the interaction. Interestingly, that relationship faded in interactions with other interlocutors. Some of these personality traits might share more or less variance with cognitive control processes and therefore, some of these processes could be mediating the association between personality traits and CS (reduced inhibition, poorer vigilance, uncertainty...) (Matthews, 2009; Mushtaq et al., 2011). A similar explanation might hold for the increased amount of CS with friends, as this is normally a more informal environment and cognitive control and performance monitoring might be reduced.

Since CS is a defining form of bi- and multilingual behaviour (Grosjean, 2010; Li Wei, 2007, 2013), it comes as no surprise that the more multilingual a person is, the greater the use of CS. It also shows that CS is linked to high levels of proficiency in different languages. This confirms the finding reported in Poplack (1980), Ritchie and Bhatia (2013), and Dewaele (2010) that a significant positive relationship exists between levels of self-perceived competence in the L2 and L3 and self-reported frequency of CS. CS is therefore ‘not an indication of a deficit in the LX but on the contrary a characteristic of participants who feel proficient in their LXs’ (2010, p. 201).

To conclude, it is worth discussing the strengths and weaknesses of our research design. We relied on participants’ self-reports concerning CS, which is not the same as actual CS by participants. We have argued earlier that multilinguals are capable of making generalisations about their CS behaviour in specific situations, and that our data have sufficient validity to investigate inter- and intra-individual variation. We feel that a research design that would rely on frequencies of actual CS would be dogged by a different and even more serious set of problems: firstly, participants would have to be wired up and masses of data would have to be recorded in a variety of settings (this would be ethical headache and a practical one: how to find researchers mastering all the languages to code the data?). While this is not impossible, the sheer amount of imposition on the participants and the amount of data to be analysed would inevitably limit the sample size to not more than a handful. Such a small sample would be insufficient for any statistical analysis into individual differences, which was the aim of the present study.

We are aware that some bi- or multilinguals with a negative perception of CS could have underestimated the frequency of their CS. There are three points to counter this criticism: 1) we reported that our participants had a generally positive attitude towards CS (Dewaele & Li Wei, 2013b); 2) the size of our sample is large enough to iron out tainted feedback from some individuals; 3) the anonymous nature of the questionnaire meant that participants could no gain anything by choosing what they perceived to be the socially desirable option.

The main advantage of the data collection via an on-line questionnaire is that it allowed us to tap into a large and culturally diverse pool of participants, which increases the ecological validity of the resulting database (Wilson & Dewaele, 2010). Our aim was to find general relationships between independent and self-reported dependent variables. The heterogeneous nature of our sample in terms of language profiles is a strength but it could also been seen as a limitation. The patterns that emerged form our data could only be described as ‘generic’. We are fully aware that in some regions of the world, the particular sociolinguistic and sociohistorical context determines language use and CS, and
the local attitudes towards these languages and CS, which could result in subtly different patterns of relationships.

Another limitation is the absence of information on whether the participants had access to CS with friends, family and co-workers. For some participants, the opportunity to CS might only exist within the family, or within the workplace. If a participant reports never to CS with colleagues, we cannot know whether this is out of choice or simply a necessity (because the colleagues do not share all the languages of the speaker).

A final limitation is that we did not differentiate between self-reported contextual switching and unintended language switching in our questionnaire. We were therefore unable to link our findings to current studies on cognitive control and bilingualism (Soveri et al., 2011). Festman (2012), for example, reported that among German L1-Russian L2 late multilinguals, the non-switchers exhibited stronger language control abilities than the habitual switchers (p. 11). Finally, it has been pointed out that frequent CS could lead to higher levels of cognitive control in bilinguals (Yim & Bialystok, 2012, p. 882). This is a fascinating perspective because it suggests that CS itself could be a source of individual differences.

Future research could focus on the relationship between same independent variables and CS in recordings of actual conversations. A more qualitative approach with a smaller number of participants could also allow an exploration of possible causes behind the patterns we uncovered. It would be particularly useful to select participants from a wide range of educational backgrounds.

**Conclusion**

The present study is the first systematic investigation of inter- and intra-individual variation in self-reported CS of a very large sample of adult multilinguals from across the world. We are aware of the need to be prudent in interpreting the present findings as we used a non-probability sampling process, and our sample contains different types of multilinguals. The statistical analyses showed that a range of both inter- and intra-individual variables are linked to the amount of self-reported CS. The intra-individual variable, namely the type of interlocutor, turned out to have a strong effect on the amount of self-reported CS: multilinguals report significantly more CS in interactions with friends, followed by interactions with colleagues and family members, and report using significantly less CS in interactions with strangers. This confirms earlier research findings (Dewaele, 2010; Resnik, 2012).

Interindividual variation in self-reported CS is also considerable. The environment in which individuals grow up and live is also linked to their self-reported frequency of CS. Interestingly, CS is clearly a phenomenon linked to higher levels of multilingualism and early bi- and multilingualism (low age of onset of acquisition, many languages known, high levels of total proficiency in the languages), again confirming earlier research (Dewaele, 2010; Poplack, 1980). Some sociobiographical variables were also found to be linked with self-reported frequency of CS: female participants reported using significantly more CS in interactions with known interlocutors. The effects of age and level of education were less clear. Finally, the present study is the first to uncover a significant relationship between two personality traits, Extraversion and Cognitive Empathy, and self-reported frequency of CS. Extraverts and participants with high levels of Cognitive Empathy report to engage more frequently in CS.

To sum up, the frequency of self-reported CS depends not just on situational, complex sociobiographical and environmental factors, but it is also mediated by the personality of the multilingual.

**Acknowledgment**

We would like to thank our friends and colleagues who forwarded the call for participation in this study, as well as all those who participated. Many thanks also to the anonymous reviewers for their excellent comments.
References


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1 Extraversion, neuroticism, openness, conscientiousness, and agreeableness.
2 Extraversion, neuroticism and psychoticism.
3 We do not report the smaller groups.
4 Dewaele and Li Wei (2013b) found that the sample of participants on which the present study is based had a mean score of 3.85 on a 5-point scale measuring positive attitudes towards CS.
5 For a more detailed discussion of this instrument, see Dewaele and Li Wei, 2013a.
6 For a more detailed discussion of this instrument, see Dewaele and Li Wei, 2012.
7 We are aware that this is an abstraction, as some may use more CS with specific individuals, but we feel that the categories reflect typical behaviour within particular language domains (Grosjean, 2010). Moreover, personality questionnaires ask even more decontextualised questions: the item ‘are you a talkative person?’ does not specify the presence of particular (types of) interlocutors or situations. Personality psychologists do not see this as posing a problem.
8 Unfortunately there is no non-parametric equivalent of multiple regression analysis which would have allowed us to measure the relative impact of all independent variables together.
9 Our sample had too few adolescents and older people (only 8 aged over 70) to carry out a comparison. We also suspect that self-reported data lack the degree of accuracy to investigate this phenomenon.