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Developmental Trajectories of Infants with Multiplex Family Risk for Autism: A Baby Siblings

Research Consortium Study

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### Key Points

**Question:** How does the development of infants with multiplex and single-incidence family risk for autism spectrum disorder (ASD) differ?

**Findings:** In this prospective, longitudinal study that included 445 children with multiplex ( $n=80$ ) or single-incidence ( $n=355$ ) family risk, 68% of children from multiplex families vs. 43% of those from single-incidence families had ASD or atypical development at outcome. Non-ASD children did not differ in ASD symptoms based on family risk status, but multiplex status was associated with lower cognitive abilities by age 3.

**Meaning:** Infants with a multiplex family history of ASD should be monitored early and often and referred for early intervention services at the first sign of concern.

## Abstract

**Importance:** Autism spectrum disorder (ASD) is a neurodevelopmental disorder associated with different genetic etiologies. Prospective examination of familial-risk infants informs understanding of developmental trajectories preceding ASD diagnosis, potentially improving early detection.

**Objective:** Compare outcomes and trajectories associated with varying familial risk for ASD across first 3 years of life.

**Design and Setting:** Longitudinal, prospective observational study. Data from 11 sites in Baby Siblings Research Consortium (BSRC) database included. Data collected between 2003-2015. Infants followed for 3 years. Analyses conducted in 2018.

**Participants:** Of initial 1,008 infants from BSRC database, 573 removed due to missing necessary data, diagnostic discrepancies, or only one older sibling. 435 younger siblings of children with ASD included; 355 from *single-incidence* families (1 sibling with ASD and 1+ sibling without ASD) and 80 from *multiplex* families (2+ siblings with ASD). No group differences in major demographics.

**Exposure:** Number of ASD-siblings.

**Main Outcomes and Measures:** Outcomes included ASD symptoms, cognitive abilities, and adaptive skills. Diagnosis (*ASD/no-ASD*) given at 36-month outcome. No-ASD group classified as *atypical* (developmental delays and/or social-communication concerns) or *typical* for some analyses. Generalized linear mixed models examined developmental trajectories by ASD outcome and familial-risk group.

**Results:** In the 435 analyzed participants (age range at outcome: 32-43 months; 57% male), children from multiplex families were more likely than those from single-incidence families to

be classified as ASD (36% vs. 16%,  $p < .001$ ) and less likely as typical (33% vs. 57%,  $p < .001$ ), with similar rates of atypical classifications (31% vs. 27%,  $p = .49$ ). No differences in ASD symptoms between multiplex and single-incidence groups, after controlling for ASD outcome ( $p = .18$ ). During infancy, differences in cognitive and adaptive abilities observed based upon ASD outcome in single-incidence group only ( $p < .001-.04$ ). At 36 months, multiplex/no-ASD group had lower cognitive abilities than single-incidence/no-ASD group ( $p = .02$ ), and multiplex had lower adaptive abilities than single-incidence, after controlling for ASD outcome ( $p = .02$ ).

**Conclusions and Relevance:** Infants with a multiplex family history of ASD should be monitored early and often and referred for early intervention at the first sign of concern. Direct examination of genetic contributions to neurodevelopmental phenotypes in infants with familial risk for ASD is needed.

Keywords: Multiplex; familial risk; autism

## Introduction

Autism spectrum disorder (ASD) is a neurodevelopmental disorder characterized by deficits in social communication and the presence of restricted and repetitive patterns of behavior<sup>1</sup>. Recent estimates indicate an ASD prevalence of 1 in 59 children and a typical age of diagnosis of 4 years old<sup>2</sup>. Converging evidence suggests that there are multiple genetic pathways to ASD<sup>3,4</sup>. One genetic risk group that has been studied widely includes infants with at least one older sibling with ASD (*familial-risk*). Prospective studies of these infants have helped to characterize the early emergence of developmental differences associated with later ASD diagnosis, with the identification of these early signs ultimately improving early screening and intervention efforts<sup>5</sup>.

To overcome the challenge of small sample sizes and to facilitate scientific collaboration in prospective studies of familial-risk infants, the Baby Siblings Research Consortium (BSRC) was formed. BSRC researchers have combined data from common measures across many sites to answer clinically-relevant questions about early manifestations of ASD. BSRC research indicates that nearly 20% of familial-risk infants will meet criteria for ASD at age 3<sup>6,7</sup> and another approximately 20% will show other developmental atypicalities (e.g., developmental delays, subclinical ASD symptoms)<sup>8,9</sup>. One key question arising from these prospective studies is whether neurodevelopmental outcomes vary based on genetic risk, with variability in risk defined by the number of siblings with ASD. *Multiplex* ASD (2+ ASD-siblings) is more commonly associated with the additive risk of common genetic variants and inherited copy number variants (CNVs)<sup>10,11</sup>, while *single-incidence* ASD (one ASD-sibling) is more often caused by rare de novo CNVs and mutations<sup>11</sup>.

Prior BSRC studies have shown that 60% of male and 30% of female children with

multiplex family risk have ASD compared to nearly 30% of male and 10% of female children with single-incidence family risk<sup>6</sup>. Profile analyses of these infants indicated that multiplex status is associated with decreased cognitive scores, but no difference in ASD symptoms<sup>6</sup>. Family-based studies have demonstrated that the non-ASD siblings of children with ASD from multiplex families have a higher level of subclinical ASD symptoms<sup>12,13</sup>, while the degree to which cognitive abilities differ among non-ASD siblings based on familial-risk status is less clear<sup>14</sup>. No studies have analyzed differences in developmental trajectories in infancy based on multiplex versus single-incidence status.

Using the BSRC database, we comprehensively examined categorical distinctions and developmental trajectories in social-communication, cognitive, and adaptive skills associated with different levels of familial risk across the first 3 years of life. We sought to answer three primary questions: (1) How do rates of typical, atypical (non-ASD), and ASD outcomes differ between infants from multiplex and single-incidence families? (2) When and how do developmental trajectories of ASD symptoms, cognitive ability, and adaptive skills across the first 3 years diverge based on familial-risk status and ASD diagnostic outcome? (3) For children without ASD, how do the phenotypic profiles differ at 3-year outcome based on familial-risk status? We expected greater impairment in infants from multiplex families versus single-incidence families, with higher rates of ASD overall, as well as lower developmental and adaptive abilities, and higher subclinical ASD symptoms in non-ASD children. Results of these analyses can help guide clinicians in earlier and more informed developmental screening and monitoring of infants from multiplex families.

## **Method**

### **Participants**



Out of an initial sample of 1,008 infants from the BSRC database, 435 younger siblings of children with ASD who were enrolled in longitudinal studies across 11 BSRC sites met inclusion criteria (partially overlapping with previous BSRC samples<sup>6,7</sup>). Children in the *multiplex* group had 2+ older siblings with ASD. Unlike previous BSRC studies<sup>6,7</sup>, children in the *single-incidence* group had a single older sibling with ASD and 1+ older sibling(s) without ASD. Confirmation of older sibling diagnoses and time points varied by study site.

Participants were removed due to missing required outcome ( $n=110$ ) or older sibling data ( $n=8$ ), discrepancies between ADOS score and diagnosis ( $n=15$ ), and having only one older sibling ( $n=404$ ) or multiple siblings from the same family ( $n=36$ ). When multiple children from a family participated, only the youngest child was included to maximize information on older siblings.

Groups were comparable with regard to demographic characteristics (Table 1). The multiplex group had larger families than the single-incidence group. IRB approval and written informed consent for all participants was obtained within each study site.

## Measures

*ASD symptoms* were measured at 18, 24, and 36 months of age using the Autism Diagnostic Observation Schedule (ADOS)<sup>15</sup>, an observational measure of social-communication and repetitive behaviors. The ADOS yields a Calibrated Severity Score (CSS) ranging from 1-10<sup>16,17</sup>. The CSS-Overall score was used in longitudinal analyses. The Social Affect (SA) and Restricted, Repetitive Behavior (RRB) subscale scores were examined in outcome analyses. The Autism Diagnostic Interview-Revised (ADI-R)<sup>18</sup>, a parent interview, was collected at 36 months in a subset of infants and used as a secondary indicator of ASD symptoms in outcome analyses.

*Cognitive abilities* were measured at 6, 9, 12, 15, 18, 24, and 36 months, using the

Mullen Scales of Early Learning (MSEL)<sup>19</sup>. The MSEL examines Visual Reception, Fine Motor, Receptive Language, and Expressive Language, which yield *t*-scores ( $M=50$ ,  $SD=10$ ). An Early Learning Composite (ELC) is also calculated, yielding a standard score ( $M=100$ ,  $SD=15$ ) representing a child's overall cognitive ability relative to peers. The ELC was used in longitudinal analyses. Subscale scores were analyzed in outcome analyses.

*Adaptive skills* were assessed at 6, 9, 12, 15, 18, 24, and 36 months in a subset of infants using the Vineland Adaptive Behavior Scales-Second Edition (Vineland-II)<sup>20</sup>, a parent-report measure. The Vineland-II assesses Communication, Daily Living Skills, Socialization, and Motor Skills, which produce standard scores. The Adaptive Behavior Composite (ABC) is computed from the first three domains, yielding a standard score representing an individual's overall adaptive ability relative to peers. The ABC was utilized in longitudinal analyses. Subscale scores were examined in outcome analyses.

*Clinical outcomes* were determined following the 36-month assessment. Children were classified as *ASD* ( $n=86$ ; vs. *no-ASD*,  $n=349$ ) if they had a clinical best estimate diagnosis of ASD by expert clinicians *and* an ADOS score at or above the clinical threshold ( $CSS \geq 4$ )<sup>7</sup>. For categorical analyses only, the no-ASD group was split into a *typical* ( $n=227$ ; MSEL ELC  $\geq 85$  and ADOS CSS  $< 3$ ) and *atypical* group ( $n=122$ ; MSEL ELC  $< 85$  and/or ADOS CSS  $\geq 3$ )<sup>8,21</sup>. Within the atypical group, 25.9% fell into this group due to lower cognitive scores, 64.7% due to elevated ADOS scores, and 9.5% due to both factors (Table 1).

## **Statistical Analyses**

Longitudinal trajectories of primary outcome variables (ADOS CSS-Overall, MSEL ELC, Vineland-II ABC) were modeled using generalized linear mixed models (GLMM) with main effects of ASD outcome (ASD vs. no-ASD), familial-risk status (multiplex vs. single-

incidence), and time, along with their two-way and three-way interactions. Subject-specific and site-specific random intercepts were included to model dependency due to repeated measures within subjects and sites. MSEL and Vineland-II scores were modeled employing an identity-link, while ADOS scores were modeled using a negative binomial GLMM with a log-link. Time was modeled as a class variable for ADOS (at 18, 24, and 36 months), with a broken-line model allowing for a slope change at 18 months for MSEL and linearly for Vineland-II (where a slope change at 18 months was non-significant). Two-way and three-way interactions between ASD outcome, familial-risk status, and time were found significant in models for Vineland-II and MSEL; however, the final GLMM for ADOS only contained the significant two-way interaction between ASD outcome and time.

According to the interactions found significant and our hypotheses, we conducted 6 contrasts for MSEL and Vineland-II data at pre-selected time points to evaluate group mean differences between: (1) ASD and no-ASD single-incidence, (2) ASD and no-ASD multiplex, (3) no-ASD multiplex and single-incidence, (4) ASD multiplex and single-incidence, (5) (ASD multiplex – no-ASD multiplex) and (ASD single-incidence – no-ASD single-incidence), (6) multiplex and single-incidence. Contrasts were conducted at 6, 12, 24, and 36 months for MSEL and 12, 24 and 36 months for Vineland-II (time points with most observations). For the final ADOS model, we conducted contrasts between: 1) ASD and no-ASD groups at 18, 24, and 36 months, and 2) multiplex and single-incidence groups (ages collapsed). We used false discovery rate (FDR)<sup>22</sup> at .05 to adjust for multiple comparisons (46 contrasts).

GLMMs account for correlations between repeated measures within subjects, allowing for fixed and time-varying covariates and automatically handling missing data, thereby

producing unbiased estimates as long as observations are missing at random. Accordingly, all available observations from each subject were utilized in modeling via GLMM.

## Results

### 36-month outcome classifications based on familial-risk status

Outcome classifications significantly differed based on familial-risk status,  $\chi^2(2, N=435)=21.10, p<.001$ . The multiplex group was more likely than the single-incidence group to be classified as ASD (36.3% vs. 16.1%),  $p<.001$ , less likely to be classified as typical (32.5% vs. 56.6%),  $p<.001$ , and had similar levels of atypical classifications (31.3% vs. 27.3%),  $p=.49$ .

### Developmental trajectories based on familial-risk status and ASD outcome

Results from the final GLMMs are summarized below. See Figure 1 for depictions of modeled developmental trajectories, Table 2 for detailed sample size information, and Table 3 for contrast results (eFigure1 presents raw trajectories).

**ASD symptoms.** ASD symptoms differed between the ASD and no-ASD groups, regardless of familial-risk status, at 18, 24, and 36 months. Children with ASD outcomes showed higher levels of ASD symptoms than children without ASD beginning at 18 months. No differences in ASD symptoms were found between the multiplex and single-incidence groups, after controlling for ASD outcome.

**Cognitive abilities.** Within the single-incidence group, children with ASD outcomes had lower cognitive abilities than no-ASD children at 6, 12, 24, and 36 months. In the multiplex group, the ASD and no-ASD groups did not differ at 6 or 12 months; instead, differences emerged at 24 months, with the ASD group demonstrating lower cognitive abilities than the no-ASD group at 24 and 36 months. Within the no-ASD group, the multiplex group had lower cognitive abilities than the single-incidence group at 36 months; cognitive abilities did not differ

based on familial-risk status among no-ASD children at earlier ages. In the ASD group, cognitive abilities did not differ between multiplex and single-incidence groups. There was, however, an overall difference in cognitive abilities between multiplex and single-incidence groups (ASD + no-ASD contrast) at 36 months. Finally, the difference in cognitive abilities among ASD and no-ASD children differed between the multiplex and single-incidence groups (ASD – no-ASD contrast) at 6 months. As depicted in Figure 1b, children with ASD outcomes had lower cognitive abilities than those without ASD within the single-incidence group at 6 months, while multiplex children had similar abilities at this age regardless of ASD outcome.

***Adaptive skills.*** Within the single-incidence group, children with ASD outcomes had lower adaptive abilities than no-ASD children at 12, 24, and 36 months. Within the multiplex group, children with and without ASD outcomes showed similar levels of adaptive abilities at 12 months, which then diverged at 24 and 36 months. The multiplex and single-incidence groups did not, however, differ significantly within the ASD and no-ASD groups. Likewise, overall familial-risk group differences were mostly non-significant. At 36 months, there were overall differences based on familial-risk status; the multiplex group had lower adaptive abilities than the single-incidence group.

### **36-month developmental profiles based on familial-risk status in no-ASD children**

See Table 4 for descriptive information and statistical results (depicted in eFigure2). Results are reported with and without correction for multiple comparisons (13 contrasts). No-ASD children from multiplex and single-incidence groups showed similar levels of social-communication skills and RRBs on the ADOS and ADI-R, and communication, socialization, daily living, and motor skills on the Vineland-II. On the MSEL, however, the multiplex group had lower visual reception and receptive language scores than the single-incidence group; the

difference in receptive language survived FDR correction.

## **Discussion**

This longitudinal investigation indicated areas of similarity and difference associated with varying levels of familial risk for ASD.

### **Rates of ASD**

Children from multiplex families were more than twice as likely to have ASD outcomes as those from single-incidence families. While 57% of the children with only one older sibling with ASD were typically developing at age 3, only 33% of the children with multiple older siblings with ASD were typically developing at outcome. This finding highlights the first and most important clinical finding of this study: infants with a strong family history of ASD need to be monitored early and often, and should be referred for early intervention services at the first sign of concern.

### **Developmental trajectories**

Longitudinal analyses suggest that group differences over time in ASD symptoms, cognitive abilities, and adaptive skills were mainly attributable to ASD outcome rather than familial-risk status. This was particularly true for ASD symptoms, which differed only based upon ASD outcome beginning at 18 months. Within the single-incidence group, children with ASD outcomes consistently demonstrated lower cognitive abilities than children without ASD beginning at 6 months and adaptive abilities beginning at 12 months (earliest ages contrasted). Conversely, multiplex infants showed similar levels of cognitive and adaptive abilities at earlier ages, regardless of ASD outcome, and did not diverge until the second year of life. Multiplex children with ASD outcomes demonstrated a sharp decline in standard scores on measures of early cognitive and adaptive skills in the second and third years of life, reflecting slower growth

in these developmental abilities. Neuroimaging studies of familial-risk infants have identified altered trajectories of brain development in the first year, particularly in cortical surface area and neural connectivity<sup>23,24</sup>. These studies have not distinguished infants based on multiplex versus single-incidence status, but they support the hypothesis that genetic risk factors lay a foundation for early changes in brain structure and function, which may then cumulatively disturb learning and adaptive behaviors leading to difficulties making expected developmental gains. These neurobiological changes may truly precede behavior; alternatively, our standardized behavioral measures may lack sensitivity to discern subtle changes in development in the first year. Clinically, these results suggest that it may be more challenging to distinguish infants with ASD versus no-ASD behaviorally in the context of multiplex status during infancy and early toddlerhood. Further research longitudinally examining biomarkers of risk early in life is needed to determine which infants are most likely to need pre-emptive intervention in this population<sup>25,26</sup>.

### **Profile analyses**

We also detected subtle differences and remarkable similarities between multiplex and single-incidence children without ASD at outcome. Non-ASD children did not differ based on familial-risk status in their observed or parent-reported levels of ASD symptoms at age 3. This was somewhat surprising given previous research suggesting subclinical ASD symptoms in family members of individuals with ASD (i.e., broader autism phenotype), particularly families with multiple affected individuals<sup>13,27</sup>. It is possible that our ASD symptom measures, which were designed as clinical diagnostic tools, were not sensitive enough to detect subtle differences in social-communication and repetitive behavior.

We did, however, detect differences in cognitive abilities at age 3. This finding was primarily explained by differences in receptive language, and, to a lesser degree, nonverbal cognitive skills, with no differences in broadly-measured expressive language found between the non-ASD multiplex and single-incidence groups. These results are largely consistent with previous research finding deficits in verbal IQ in the unaffected siblings of multiplex but not single-incidence families<sup>14</sup>. The likely risk factors for having multiple children with ASD, such as shared genetic variation, vulnerability to genetic mutations, or complex gene-environment interactions (e.g., in utero environment) may impact brain development in a more distributed, global way, which then impacts overall development, rather than networks that are more specific to ASD. These findings speak to the need for large, collaborative efforts to examine brain development, genetics, and gene-environment interactions in at-risk infants to understand the neurobiological mechanisms underlying these differences in developmental trajectories and behavior.

### **Strengths and limitations**

Our study uniquely leveraged a rich dataset collected from multiple expert sites to prospectively examine differences associated with multiplex status and diagnostic outcome in a large cohort of infants with elevated familial risk for ASD. Although the sample size was quite large for a study of this kind, the prospective nature led to uneven and occasionally small groups disallowing investigation in some areas of interest (e.g., sex) and firm conclusions in others. For instance, the multiplex group was smaller, so comparisons within this group were less powered than those within the single-incidence group. Given the longitudinal, multi-site design, there was also some inconsistency among study sites in the ages at which different measures were collected, how older sibling diagnoses were confirmed, as well as missing data. Statistical



models that account for missing data and site differences helped to attenuate possible negative effects. The use of already-collected data across multiple sites also required us to choose common broad-based measures that, while highly clinically relevant and well-validated, may not have been sensitive enough to detect more subtle differences between the non-ASD children. Additionally, as is the case across the ASD-sibling literature<sup>28</sup>, many of the children in the sample had relatively high cognitive scores and came from predominantly Caucasian and highly-educated families, so these results may not represent the larger population of children with ASD. The most substantial limitation is the lack of genomic data in these infants, which would inform our hypotheses about genetic factors contributing to developmental differences.

### **Conclusions**

Children from multiplex families are more than twice as likely to meet criteria for ASD at age 3 than children from single-incidence families. Prospectively, single-incidence infants begin to show developmental differences based on later ASD diagnosis by 6 months of age, while multiplex infants with and without ASD outcomes do not differ until the second year of life. Among unaffected children, multiplex risk is associated with lower cognitive abilities, but similar levels of ASD symptoms. Results support the need for direct examination of genetic contributions to neurodevelopmental phenotypes in infants with multiplex and single-incidence family risk for ASD. Given their very high rates of ASD and other neurodevelopmental challenges, infants with a strong family history of ASD should be monitored early and often and referred for early intervention at the first sign of developmental concern.

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*Concept and design:* McDonald, Jeste.

*Acquisition, analysis, or interpretation of data:* All authors.

*Drafting of the manuscript:* McDonald, Jeste, Senturk.

*Critical revision of the manuscript for important intellectual content:* All authors.

*Statistical analysis:* McDonald, Senturk, Scheffler.

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## References

1. American Psychiatric Association. *Diagnostic and Statistical Manual of Mental Disorders*. 5th ed. Washington D.C.: 2013.
2. Baio J, Wiggins L, Christensen DL, et al. Prevalence of autism spectrum disorder among children aged 8 years—Autism and Developmental Disabilities Monitoring Network, 11 sites, United States, 2014. *MMWR Surveill Summ*. 2018;67(SS-6):1-23. doi: <http://dx.doi.org/10.15585/mmwr.ss6706a1>
3. Ramaswami G, Geschwind DH. Genetics of autism spectrum disorder. *Handb Clin Neurol*. 2018;147:321-329. doi: 10.1016/B978-0-444-63233-3.00021-X
4. Muhle RA, Reed HE, Stratigos KA, Veenstra-VanderWeele J. The emerging clinical neuroscience of autism spectrum disorder: A review. *JAMA Psychiatry*. 2018;75(5):514-523. doi: 10.1001/jamapsychiatry.2017.4685
5. Jones EJH, Gliga T, Bedford R, Charman T, Johnson MH. Developmental pathways to autism: A review of prospective studies of infants at risk. *Neurosci Biobehav Rev*. 2014;39:1-33. doi: 10.1016/j.neubiorev.2013.12.001
6. Messinger DS, Young GS, Webb SJ, et al. Early sex differences are not autism-specific: A Baby Siblings Research Consortium (BSRC) study. *Mol Autism*. 2015;6:32. doi: 10.1186/s13229-015-0027-y
7. Ozonoff S, Young GS, Carter A, et al. Recurrence risk for autism spectrum disorders: A Baby Siblings Research Consortium study. *Pediatrics*. 2011;128(3):e488-e495. doi: 10.1542/peds.2010-2825

8. Charman T, Young GS, Brian J, et al. Non-ASD outcomes at 36 months in siblings at familial risk for autism spectrum disorder (ASD): A baby siblings research consortium (BSRC) study. *Autism Res.* 2017;10(1):169-178. doi: 10.1002/aur.1669
9. Messinger D, Young GS, Ozonoff S, et al. Beyond autism: A baby siblings research consortium study of high-risk children at three years of age. *J Am Acad Child Adolesc Psychiatry.* 2013;52(3):300-308.e1. doi: 10.1016/j.jaac.2012.12.011
10. Klei L, Sanders SJ, Murtha MT, et al. Common genetic variants, acting additively, are a major source of risk for autism. *Mol autism.* 2012;3(1):9. doi: 10.1186/2040-2392-3-9
11. Leppa VM, Kravitz SN, Martkin CL, et al. Rare inherited and de novo CNVs reveal complex contributions to ASD risk in multiplex families. *Am J Hum Genet.* 2016;99(3):540-554. doi: 10.1016/j.ajhg.2016.06.036
12. Frazier TW, Youngstrom EA, Hardan AY, Georgiades S, Constantino JN, Eng C. Quantitative autism symptom patterns recapitulate differential mechanisms of genetic transmission in single and multiple incidence families. *Mol autism.* 2015;6:58. doi: 10.1186/s13229-015-0050-z
13. Gerds JA, Bernier R, Dawson G, Estes A. The broader autism phenotype in simplex and multiplex families. *J Autism Dev Disord.* 2013;43(7):1597-605. doi: 10.1007/s10803-012-1706-6
14. Oerlemans AM, Hartman CA, Franke B, Buitelaar JK, Rommelse NN. Does the cognitive architecture of simplex and multiplex ASD families differ? *J Autism Dev Disord.* 2016;46(2):489-501. doi: 10.1007/s10803-015-2572-9

15. Lord C, Risi S, Lambrecht L, et al. The autism diagnostic observation schedule-generic: a standard measure of social and communication deficits associated with the spectrum of autism. *J Autism Dev Disord.* 2000;30(3):205-223.
16. Gotham K, Pickles A, Lord C. Standardizing ADOS scores for a measure of severity in autism spectrum disorders. *J Autism Dev Disord.* 2009;39(5):693-705. doi: 10.1007/s10803-008-0674-3
17. Hus V, Gotham K, Lord C. Standardizing ADOS domain scores: separating severity of social affect and restricted and repetitive behaviors. *J Autism Dev Disord.* 2014;44(10):2400-2412. doi: 10.1007/s10803-012-1719-1
18. Rutter M, Le Couteur A, Lord C. *ADI-R Autism Diagnostic Interview Revised. Manual.* Los Angeles: Western Psychological Services. 2013.
19. Mullen EM. *Mullen Scales Of Early Learning.* CirclePines, MN: American Guidance Service. 1995.
20. Sparrow SS, Cicchetti VD, Balla AD. *Vineland Adaptive Behavior Scales.* 2nd Edition. American Guidance Service; Circle Pines, MN: 2005.
21. Chawarska K, Shic F, Macari S, et al. 18-month predictors of later outcomes in younger siblings of children with autism spectrum disorder: A Baby Siblings Research Consortium study. *J Am Acad Child Adolesc Psychiatry.* 2014;53(12):1317-1327.e1311. doi: 10.1016/j.jaac.2014.09.015
22. Benjamini Y, Hochberg Y. Controlling the false discovery rate: A practical and powerful approach to multiple testing. *J R Stat Soc Series B Stat Methodol.* 1995;57(1):289-300.
23. Hazlett HC, Gu H, Munsell BC, et al. Early brain development in infants at high risk for autism spectrum disorder. *Nature.* 2017;542(7641):348-351. doi: 10.1038/nature21369

24. Lewis JD, Evans AC, Pruett JR, Jr., et al. The emergence of network inefficiencies in infants with autism spectrum disorder. *Biol Psychiatry*. 2017;82(3):176-185. doi: 10.1016/j.biopsych.2017.03.006
25. Green J, Pickles A, Pasco G, et al. Randomised trial of a parent-mediated intervention for infants at high risk for autism: longitudinal outcomes to age 3 years. *J Child Psychol Psychiatry*. 2017;58(12):1330-1340. doi: 10.1111/jcpp.12728
26. Jones EJH, Dawson G, Kelly J, Estes A, Webb SJ. Parent-delivered early intervention in infants at risk for ASD: Effects on electrophysiological and habituation measures of social attention. *Autism Res*. 2017;10(5):961-972. doi: 10.1002/aur.1754
27. Schwichtenberg AJ, Young GS, Sigman M, Hutman T, Ozonoff S. Can family affectedness inform infant sibling outcomes of autism spectrum disorders? *J Child Psychol Psychiatry*. 2010;51(9):1021-30. doi: 10.1111/j.1469-7610.2010.02267.x
28. Sacrey LR, Zwaigenbaum L, Szatmari P, et al. Brief report: Characteristics of preschool children with ASD vary by ascertainment. *J Autism Dev Disord*. 2017;47(5):1542-1550. doi: 10.1007/s10803-017-3062-z

**Table 1.** Participant information by familial-risk and outcome group

<b>Variable</b>	<b>Single-incidence (n=355)</b>	<b>Multiplex (n=80)</b>	<b><i>p</i></b>
Sex (male) <i>n</i> (%)	200 (56.3)	46 (57.5)	.85
Race/ethnicity (non-Caucasian) <i>n</i> (%)	86 (24.3)	14 (17.5)	.36
Maternal education (college or higher) <i>n</i> (%)	233 (65.7)	53 (66.7)	.92
Maternal age at birth (years) <i>M</i> ( <i>SD</i> )	34.86 (4.82)	34.57 (4.87)	.69
Paternal age at birth (years) <i>M</i> ( <i>SD</i> )	37.37 (5.81)	37.35 (6.26)	.98
Number of children in family <i>M</i> ( <i>SD</i> )	3.46 (.79)	3.83 (1.35)	.02
Number of ASD siblings <i>M</i> ( <i>SD</i> )	1.00 (0)	2.13 (.44)	<.001
Age first seen (months) <i>M</i> ( <i>SD</i> )	6.90 (4.24)	7.23 (4.14)	.53
Age at outcome (months) <i>M</i> ( <i>SD</i> )	37.28 (1.63)	37.48 (1.92)	.34
<b>36-month outcome variable*</b>	<b>Typical (n=227)</b>	<b>Atypical (n=122)</b>	<b>ASD (n=86)</b>
<b><i>ASD symptoms (ADOS) M (SD)</i></b>			
Overall	1.30 (.46)	3.52 (1.77)	6.85 (1.78)
Social Affect	1.64 (.77)	3.85 (2.08)	6.67 (1.83)
RRB	3.10 (2.37)	5.07 (2.45)	7.58 (1.91)
<b><i>Cognitive abilities (MSEL) M (SD)</i></b>			
Early Learning Composite	110.49 (14.37)	96.90 (19.84)	81.06 (24.42)
Visual Reception	61.25 (10.31)	53.14 (15.14)	42.54 (18.84)
Fine Motor	52.90 (12.65)	44.73 (12.81)	36.12 (12.75)
Receptive Language	52.77 (8.84)	46.61 (9.86)	38.52 (14.31)
Expressive Language	54.14 (7.82)	47.84 (10.94)	39.81 (13.56)
<b><i>Adaptive skills (Vineland-II) M (SD)</i></b>			
Adaptive Behavior Composite	94.95 (12.11)	90.37 (13.84)	79.14 (13.73)
Communication	100.90 (13.18)	96.12 (14.16)	85.22 (16.32)
Daily Living Skills	94.95 (11.41)	89.95 (15.90)	80.55 (14.04)
Socialization	97.66 (12.68)	92.97 (12.85)	79.59 (13.21)
Motor Skills	94.88 (12.06)	92.00 (12.71)	84.83 (12.78)



*Note.* Group differences between categorical variables assessed using chi-square tests, and continuous variables using t-tests. ADOS calibrated severity score (1-10), Vineland-II composite and domain standard scores ( $M=100$ ,  $SD=15$ ), MSEL composite standard score ( $M=100$ ,  $SD=15$ ) and subscale t-scores ( $M=50$ ,  $SD=10$ ) presented. \*Groups differed on all outcome variables.

**Table 2.** Number of participants with data by age, measure, and group status

<b>Age (mos)</b>	<b>Single-incidence/ No-ASD</b>	<b>Single-incidence/ ASD</b>	<b>Multiplex/ No-ASD</b>	<b>Multiplex/ ASD</b>
Total in sample	298	57	51	29
<b><i>ADOS</i></b>				
18	214	39	36	23
24	260	54	44	27
36	298	57	51	29
<b><i>MSEL</i></b>				
6	175	29	28	13
9	48	11	14	5
12	249	38	41	23
15	57	16	11	7
18	113	19	18	15
24	266	49	48	27
36	294	56	49	27
<b><i>Vineland-II</i></b>				
6	38	9	5	2
9	38	9	4	3
12	148	26	21	14
15	42	13	7	3
18	177	28	29	17
24	187	37	28	18
36	210	39	30	17
<b><i>ADI-R</i></b>				
36	136	37	28	15

*Note.* GLMM models used all available data to inform estimates.

**Table 3.** GLMM contrast results

<b>Group</b>	<b>Contrast</b>	<b>Age (mos)</b>	<b>Estimate (SE)</b>	<b>df</b>	<b>t</b>	<b>p</b>	<b>f</b>
<i>Observed ASD symptoms (ADOS)</i>							
-	ASD vs. No-ASD	18	.73 (.08)	693	8.95	<.001	.34
-	ASD vs. No-ASD	24	.78 (.07)	693	10.95	<.001	.42
-	ASD vs. No-ASD	36	1.16 (.07)	693	17.94	<.001	.68
-	Multiplex vs. Single	-	.09 (.06)	693	1.67	.18	.06
<b>Group</b>	<b>Contrast</b>	<b>Age (mos)</b>	<b>Estimate (SE)</b>	<b>df</b>	<b>t</b>	<b>p</b>	<b>f</b>
<i>Cognitive abilities (MSEL)</i>							
Single-incidence	ASD vs. No-ASD	6	-6.52 (2.67)	1304	-2.45	.04	.07
Single-incidence	ASD vs. No-ASD	12	-10.15 (2.02)	1304	-5.04	<.001	.14
Single-incidence	ASD vs. No-ASD	24	-17.53 (2.04)	1304	-8.58	<.001	.24
Single-incidence	ASD vs. No-ASD	36	-25.01 (2.35)	1304	-10.65	<.001	.30
Multiplex	ASD vs. No-ASD	6	5.67 (4.30)	1304	1.32	.29	.04
Multiplex	ASD vs. No-ASD	12	-5.89 (3.17)	1304	-1.86	.13	.05
Multiplex	ASD vs. No-ASD	24	-18.84 (3.18)	1304	-5.92	<.001	.16
Multiplex	ASD vs. No-ASD	36	-21.61 (3.81)	1304	-5.68	<.001	.16
No-ASD	Multiplex vs. Single	6	-2.48 (2.67)	1304	-.93	.44	.03
No-ASD	Multiplex vs. Single	12	-.65 (2.05)	1304	-.32	.79	.01
No-ASD	Multiplex vs. Single	24	-1.56 (2.11)	1304	-.74	.51	.02
No-ASD	Multiplex vs. Single	36	-7.05 (2.47)	1304	-2.85	.02	.08
ASD	Multiplex vs. Single	6	9.71 (4.31)	1304	2.26	.05	.06
ASD	Multiplex vs. Single	12	3.61 (3.16)	1304	1.14	.36	.03

ASD	Multiplex vs. Single	24	-2.88 (3.15)	1304	-.91	.44	.03
ASD	Multiplex vs. Single	36	-3.65 (3.74)	1304	-.98	.43	.03
ASD – No-ASD	Multiplex vs. Single	6	12.20 (5.06)	1304	2.41	.04	.07
ASD – No-ASD	Multiplex vs. Single	12	4.26 (3.76)	1304	1.13	.36	.03
ASD – No-ASD	Multiplex vs. Single	24	-1.32 (3.79)	1304	-.35	.78	.01
ASD – No-ASD	Multiplex vs. Single	36	3.40 (4.48)	1304	.76	.50	.02
ASD + No-ASD	Multiplex vs. Single	6	3.61 (2.54)	1304	1.43	.26	.04
ASD + No-ASD	Multiplex vs. Single	12	1.48 (1.89)	1304	.79	.50	.02
ASD + No-ASD	Multiplex vs. Single	24	-2.22 (1.90)	1304	-1.17	.36	.03
ASD + No-ASD	Multiplex vs. Single	36	-5.35 (2.25)	1304	-2.38	<.05	.07

<b>Group</b>	<b>Contrast</b>	<b>Age (mos)</b>	<b>Estimate (SE)</b>	<b>df</b>	<b>t</b>	<b>p</b>	<b>f</b>
<i>Adaptive skills (Vineland-II)</i>							
Single-incidence	ASD vs. No-ASD	12	-7.59 (1.76)	843	-4.32	<.001	.15
Single-incidence	ASD vs. No-ASD	24	-9.82 (1.52)	843	-6.43	<.001	.22
Single-incidence	ASD vs. No-ASD	36	-12.05 (1.94)	843	-6.21	<.001	.22
Multiplex	ASD vs. No-ASD	12	-2.62 (3.01)	843	-0.87	.45	.03
Multiplex	ASD vs. No-ASD	24	-9.38 (2.50)	843	-3.75	.001	.13
Multiplex	ASD vs. No-ASD	36	-16.14 (3.32)	843	-4.87	<.001	.17
No-ASD	Multiplex vs. Single	12	-4.14 (1.93)	843	-2.14	.07	.07
No-ASD	Multiplex vs. Single	24	-3.67 (1.61)	843	-2.28	.05	.08
No-ASD	Multiplex vs. Single	36	-3.20 (2.12)	843	-1.51	.23	.05
ASD	Multiplex vs. Single	12	.82 (2.90)	843	0.28	.79	.01
ASD	Multiplex vs. Single	24	-3.23 (2.45)	843	-1.32	.29	.05

ASD	Multiplex vs. Single	36	-7.29 (3.20)	843	-2.28	.05	.08
ASD – No-ASD	Multiplex vs. Single	12	4.96 (3.49)	843	1.42	.26	.05
ASD – No-ASD	Multiplex vs. Single	24	.44 (2.93)	843	.15	.88	.01
ASD – No-ASD	Multiplex vs. Single	36	-4.09 (3.84)	843	-1.07	.39	.04
ASD + No-ASD	Multiplex vs. Single	12	-1.65 (1.74)	843	-.95	.46	.03
ASD + No-ASD	Multiplex vs. Single	24	-2.56 (1.51)	843	-1.69	.18	.06
ASD + No-ASD	Multiplex vs. Single	36	-5.25 (1.92)	843	-2.73	.02	.08

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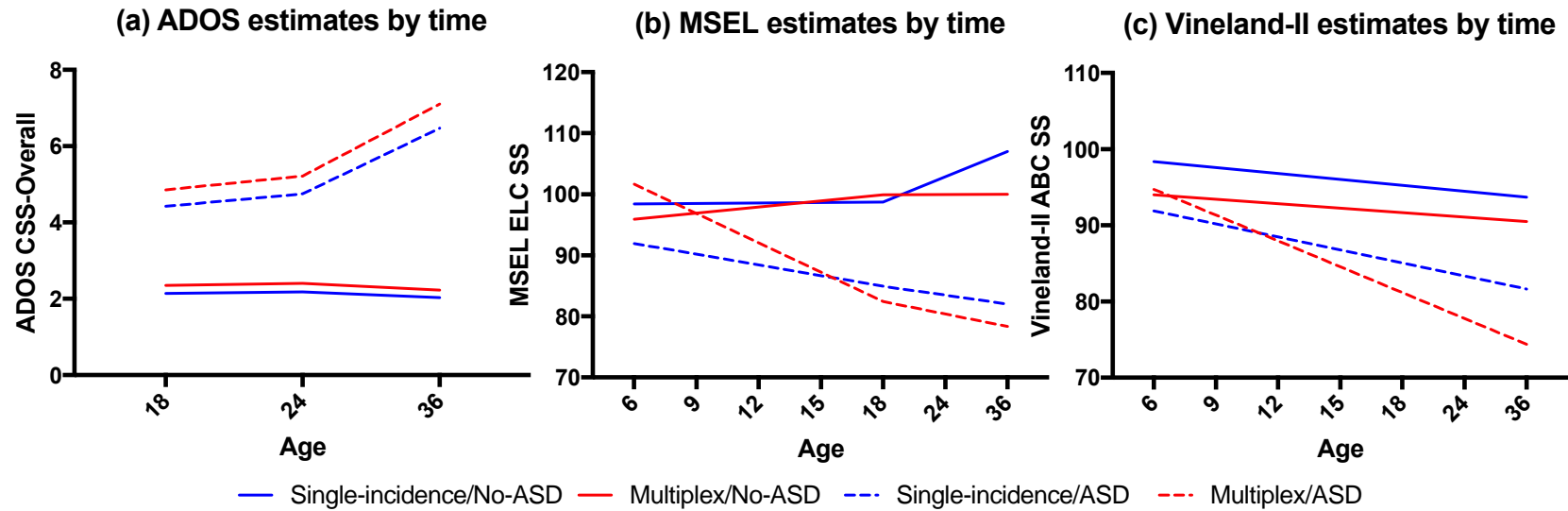
*Note.* Reported  $p$ -values are FDR corrected.

**Table 4.** Detailed comparison of 36-month outcome data across familial-risk groups in no-ASD children

<b>Variable</b> <i>M (SD)</i>	<b>Single-incidence</b>	<b>Multiplex</b>	<i>p<sup>raw</sup></i>	<i>p<sup>FDR</sup></i>	<i>d</i>
<b><i>Cognitive abilities (MSEL)</i></b>					
Visual Reception	59.14 (12.26)	54.70 (14.85)	.05	.22	.33
Fine Motor	50.66 (13.03)	46.96 (14.32)	.07	.22	.27
Receptive Language	51.80 (9.91)	47.61 (8.80)	.003	.04	.45
Expressive Language	52.28 (9.38)	50.39 (9.89)	.20	.29	.20
<b><i>Observed ASD symptoms (ADOS)</i></b>					
Social Affect	2.38 (1.74)	2.63 (1.68)	.35	.41	.15
RRB	3.68 (2.58)	4.41 (2.48)	.06	.22	.29
<b><i>Reported ASD symptoms (ADI-R)</i></b>					
Social Interaction	2.66 (2.30)	3.82 (3.84)	.13	.29	.37
Communication	2.39 (2.50)	3.29 (3.47)	.20	.29	.30
RRB	.73 (1.39)	1.29 (1.86)	.14	.29	.34
<b><i>Adaptive skills (Vineland-II)</i></b>					
Communication	99.57 (13.61)	96.03 (14.33)	.17	.29	.25
Daily Living Skills	93.47 (13.43)	90.24 (13.71)	.23	.30	.24
Socialization	96.17 (12.03)	93.85 (17.62)	.46	.50	.15
Motor Skills	93.90 (11.62)	93.07 (16.81)	.80	.80	.06

*Note.* Both uncorrected ( $p^{raw}$ ) and FDR-corrected ( $p^{FDR}$ ) *p*-values reported.

**Figure 1.** Developmental trajectories of (a) ASD symptoms, (b) cognitive abilities, and (c) adaptive skills



*Figure 1 caption.* Depiction of results from Generalized Linear Mixed Models. ADOS measured ASD symptoms, MSEL measured cognitive abilities, and Vineland-II measured adaptive skills. ADOS=Autism Diagnostic Observation Schedule. CSS=Calibrated Severity Score. MSEL=Mullen Scales of Early Learning. ELC=Early Learning Composite. SS=Standard Score. ABC=Adaptive Behavior Composite.